Virtual Branch Analysis of Symbol Error Probability for Hybrid Selection/Maximal-Ratio Combining in Rayleigh Fading

Moe Z. Win, Senior Member, IEEE, and Jack H. Winters, Fellow, IEEE

Abstract—In this paper, we derive analytical expressions for the symbol error probability (SEP) for a hybrid selection/maximalratio combining (H-S/MRC) diversity system in multipath-fading wireless environments. With H-S/MRC, L out of N diversity branches are selected and combined using maximal-ratio combining (MRC). We consider coherent detection of M-ary phase-shift keying (MPSK) and quadrature amplitude modulation (MQAM) using H-S/MRC for the case of independent Rayleigh fading with equal signal-to-noise ratio averaged over the fading. The proposed problem is made analytically tractable by transforming the ordered physical diversity branches, which are correlated, into independent and identically distributed (i.i.d.) "virtual branches," which results in a simple derivation of the SEP for arbitrary L and N. We further obtain a canonical structure for the SEP of H-S/MRC as a weighted sum of the elementary SEP's, which are the SEP's using MRC with i.i.d. diversity branches in Rayleigh fading, or equivalently the SEP's of the nondiversity (single-branch) system in Nakagami fading, whose closed-form expressions are well-known. We present numerical examples illustrating that H-S/MRC, even with $L \ll N$, can achieve performance close to that of N-branch MRC.

Index Terms—Diversity combining, error probability, fading channel, maximal ratio combining, selection diversity, virtual branch technique.

I. INTRODUCTION

T HE CAPACITY of wireless systems in a multipath environment can be increased by diversity techniques [1], such as selection diversity (SD) [2]–[4] or maximal-ratio combining (MRC) [4]. SD is the simplest form of diversity system whereby the received signal is selected from *one* out of N available diversity branches. In MRC, the received signals from *all* the diversity branches are weighted and combined to maximize the *instantaneous* signal-to-noise ratio (SNR) at the combiner output.

Though a high diversity order is possible in many situations, it may not be feasible to utilize all of the available branches. For example, a large order of antenna diversity may be obtained easily, especially at higher frequencies such as the PCS bands, using spatial separation and/or orthogonal polarizations. Even for a handset, the main diversity-order limitation is typically not the handset size (which determines the maximum number

The authors are with the Wireless Systems Research Department, AT&T Labs–Research, Middletown, NJ 07748-4801 USA (e-mail: win@research.att. com; jhw@research.att.com).

Publisher Item Identifier S 0090-6778(01)10177-7.

of antenna elements) but the power consumption and cost of the RF electronics for each diversity branch [5].

This has motivated studies [6]-[11] of diversity combining techniques that process only a *subset* of the available diversity branches with limited resources (i.e., power, RF electronics), but achieve better performance than SD. These reduced-complexity combining systems select L branches (from N available diversity branches) and combine them based on a chosen criterion. Here, we consider a hybrid selection/maximal-ratio combining (H-S/MRC) diversity system which selects the L branches with largest receive SNR at each instant, and then combines these branches to maximize the instantaneous output SNR. This potentially reduces the number of required RF chains from N to L. We assume that instantaneous channel estimation using a scanning receiver across all possible diversity branches is feasible, such as with slow fading. However, H-S/MRC also offers improvement in fast fading conditions, and our results serve as a lower bound on the symbol error probability (SEP) performance when perfect channel estimates are not available.

In [7], the bit error rate performance of H-S/MRC with L = 2and L = 3 out of N branches was analyzed, and it was pointed out that "the expressions become extremely unwieldy" for L >3. The *average* SNR of H-S/MRC was derived in [8]. In [10], a "virtual branch" technique was introduced (see also [11]) to succinctly derive the mean as well as the variance of the combiner output SNR of the H-S/MRC diversity system. Concurrent and independent work on the performance analysis of H-S/MRC can also be found in [12], where H-S/MRC is referred to as "generalized selection combining."

In this paper, we derive *exact* expressions for the SEP of a H-S/MRC diversity system with *arbitrary* L and N. We consider coherent detection of M-ary phase-shift keying (MPSK) and quadrature amplitude modulation (MQAM) for the case of independent Rayleigh fading with equal SNR averaged over the fading. The proposed problem is made analytically tractable by transforming the ordered physical diversity branches, which are correlated, into independent and identically distributed (i.i.d.) *virtual branches*.¹ We further obtain a *canonical structure* for the SEP of H-S/MRC as a weighted sum of the elementary SEP's, which are the SEP's using MRC with i.i.d. branches in Rayleigh fading, or equivalently the SEP's of a nondiversity (single-branch) system in Nakagami fading, whose closed-form

Paper approved by K.-C. Chen, the Editor for Wireless data Communications of the IEEE Communications Society. Manuscript received June 4, 1999; revised December 21, 1999. This paper was presented in part at the IEEE Global Telecommunications Conference, Rio de Janeiro, Brazil, December 1999.

¹When the average branch SNR's are not necessarily equal, it can be shown that the virtual branch technique still applies, but the virtual branches are *conditionally* independent [9].

expressions are well known. We also present numerical examples illustrating that H-S/MRC, even with $L \ll N$, can achieve performance close to that of N-branch MRC.

II. DIVERSITY COMBINING ANALYSIS

A. Virtual Branch Technique: The Key Idea

The analysis of H-S/MRC based on a chosen ordering of the branches at first appears to be complicated, since the SNR statistics of the ordered branches are *not* independent. Here, we alleviate this problem by transforming the ordered-branch variables into a new set of i.i.d. *virtual branches*, and expressing the ordered-branch SNR variables as a linear function of i.i.d. virtual branch SNR variables. The key advantage of this formulation is that it allows greater flexibility in the selection process of the ordered instantaneous SNR values, and permits the combiner output SNR to be expressed in terms of the i.i.d. virtual branch SNR variables. In this framework, the derivation of the SEP for H-S/MRC, involving the evaluation of nested *N*-fold integrals, essentially reduces to the evaluation of a single integral with *finite* limits. The well-known results for SD and MRC are shown to be special cases of our results.

B. General Theory

Let γ_i denote the instantaneous SNR of the *i*th diversity branch defined by

$$\gamma_i \stackrel{\Delta}{=} \alpha_i^2 \frac{E_{\rm s}}{N_{0i}} \tag{1}$$

where $2E_s$ is the average symbol energy, and α_i is the instantaneous fading amplitude and $2N_{0i}$ is the two-sided noise power spectral density of the *i*th branch. We model the γ_i 's as continuous random variables with (probability density function (pdf) $f_{\gamma_i}(x)$ and mean $\Gamma_i = \mathbb{E}\{\gamma_i\}$.

Let us first consider a general diversity combining (GDC) system with the instantaneous output SNR of the form²

$$\gamma_{\rm GDC} = \langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]} \rangle \tag{2}$$

where \boldsymbol{a} and $\boldsymbol{\gamma}_{[N]}$ are $N \times 1$ vectors with N denoting the number of available diversity branches. The selection vector \boldsymbol{a} is binary-valued with *i*th element $a_i \in \{0, 1\}$. The ordered vector $\boldsymbol{\gamma}_{[N]} \triangleq [\gamma_{[1]}, \gamma_{[2]}, \dots, \gamma_{[N]}]^t$, where $\{\gamma_{[i]}\}$ is the ordered set of $\{\gamma_i\}$, i.e., $\gamma_{[1]} > \gamma_{[2]} > \dots > \gamma_{[N]}$ and $(\cdot)^t$ denotes transpose. Note that GDC selectively combines the branches with instantaneous SNR $\gamma_{[i]}$ corresponding to nonzero elements $(a_i \neq 0)$ of the selection vector \boldsymbol{a} . It will be apparent later that several diversity combining schemes, including H-S/MRC, turn out to be special cases of (2). Note that the possibility of at least two equal $\gamma_{[i]}$'s is excluded, since $\gamma_{[i]} \neq \gamma_{[j]}$ almost surely for continuous random variables γ_i 's.³ For a Rayleigh fading channel, the pdf of α_i 's is given by

$$f_{\alpha_i}(r) = \frac{2r}{\Omega_i} e^{-r^2/\Omega_i}, \qquad r \ge 0 \tag{3}$$

where $\Omega_i = \mathbb{E}\{\alpha_i^2\}$. Then, the pdf of the instantaneous branch SNR is given by

$$f_{\gamma_i}(x) = \begin{cases} \frac{1}{\Gamma_i} e^{-(x/\Gamma_i)}, & 0 \le x < \infty\\ 0, & \text{otherwise} \end{cases}$$
(4)

where the mean $\Gamma_i = \mathbb{E}\{\gamma_i\} = \mathbb{E}\{\alpha_i^2\}(E_s/N_{0i}) = \Omega_i(E_s/N_{0i})$. If γ_i s are independent with equal average SNR, i.e., $\Gamma_i = \Gamma$ for i = 1, ..., N, then $f_{\gamma_i}(x) = f(x)$ for i = 1, ..., N.

The joint pdf of $\gamma_{[1]}, \gamma_{[2]}, \ldots, \gamma_{[N]}$ can be derived using the theory of "order statistics" [17] as

$$f_{\boldsymbol{\gamma}_{[N]}}\left(\left\{\gamma_{[i]}\right\}_{i=1}^{N}\right) = \begin{cases} N! f(\gamma_{[1]}) f(\gamma_{[2]}) \cdots f(\gamma_{[N]}), \\ \gamma_{[1]} > \gamma_{[2]} > \cdots > \gamma_{[N]} \\ 0, \quad \text{otherwise.} \end{cases}$$
(5)

Therefore

$$f_{\boldsymbol{\gamma}_{[N]}}\left(\left\{\gamma_{[i]}\right\}_{i=1}^{N}\right) = \begin{cases} N! \left(\frac{1}{\Gamma}\right)^{N} e^{-(1/\Gamma)\langle \mathbf{1}_{N}, \boldsymbol{\gamma}_{[N]} \rangle}, \\ \gamma_{[1]} > \gamma_{[2]} > \dots > \gamma_{[N]} > 0 \\ 0, \quad \text{otherwise} \end{cases}$$
(6)

where $\mathbf{1}_N$ denotes the vector of length N whose elements are all ones. It is important to note that the $\gamma_{[i]}$'s are *no* longer independent, even though the underlying γ_i 's are independent.

C. Symbol Error Probability for GDC Over the Channel Ensemble

The SEP for GDC in multipath-fading environment is obtained by averaging the conditional SEP over the channel ensemble. This can be accomplished by averaging the $Pr\{e|\gamma_{GDC}\}$ over the pdf of γ_{GDC} as

$$P_{e,\text{GDC}} = \mathbb{E}_{\gamma_{\text{GDC}}} \left\{ \Pr\left\{ e | \gamma_{\text{GDC}} \right\} \right\}$$
$$= \int_{0}^{\infty} \Pr\{e | \gamma\} f_{\gamma_{\text{GDC}}}(\gamma) d\gamma$$
(7)

where $\Pr\{e|\gamma_{GDC}\}$ is the *conditional* SEP, conditioned on the random variable γ_{GDC} , and $f_{\gamma_{GDC}}(\cdot)$ is the pdf of the combiner output SNR [18]–[20]. Alternatively, averaging over the channel ensemble can be accomplished, using the technique of [21], [22], by substituting the expression for γ_{GDC} directly in terms of the physical branch variables given in (2), as

$$P_{e,\text{GDC}} = \mathbb{E}_{\{\gamma_{[i]}\}} \left\{ \Pr\left\{ e | \gamma_{\text{GDC}} = \langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]} \rangle \right\} \right\}$$
$$= \int_{0}^{\infty} \int_{0}^{\gamma_{[1]}} \cdots \int_{0}^{\gamma_{[N-1]}} \Pr\left\{ e | \langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]} \rangle \right\}$$
$$\times f_{\boldsymbol{\gamma}_{[N]}} \left(\left\{ \gamma_{[i]} \right\}_{i=1}^{N} \right) d\gamma_{[N]} \cdots d\gamma_{[2]} d\gamma_{[1]}. \tag{8}$$

Since the statistics of the ordered-branches are no longer independent, the evaluation of (8) involves nested N-fold integrals, which are in general cumbersome and complicated to

²The notation $\langle \boldsymbol{x}, \boldsymbol{y} \rangle$ for $\boldsymbol{x}, \boldsymbol{y} \in \mathbb{R}^N$ is used to denote the usual inner product on \mathbb{R}^N defined by $\langle \boldsymbol{x}, \boldsymbol{y} \rangle \stackrel{\Delta}{=} \sum_{i=1}^N x_i y_i$. For a linear transformation $T: \mathbb{R}^N \to \mathbb{R}^N$, we will use the fact that $\langle \boldsymbol{x}, T \boldsymbol{y} \rangle = \langle T^i \boldsymbol{x}, \boldsymbol{y} \rangle$ [13], [14].

³In our context, the notion of "almost sure" or "almost everywhere" can be stated mathematically as: $\gamma_{[i]} \neq \gamma_{[j]}$ almost surely if and only if $\Pr\{\gamma_{[i]} = \gamma_{[j]}\} = 0$ [15], [16].

compute. This can be alleviated by transforming the instantaneous SNR of the ordered diversity branches, $\gamma_{[i]}$'s, into a new set of virtual branch instantaneous SNR's, V_n 's, using the following relation:

$$\boldsymbol{\gamma}_{[N]} = \boldsymbol{T}_{\mathrm{VB}} \boldsymbol{V}_N \tag{9}$$

where T_{VB} : $\mathbb{R}^N \to \mathbb{R}^N$ is the upper triangular virtual branch transformation matrix given by

$$\boldsymbol{T}_{\mathrm{VB}} = \begin{bmatrix} \frac{\Gamma}{1} & \frac{\Gamma}{2} & \cdots & \frac{\Gamma}{N} \\ & \frac{\Gamma}{2} & \cdots & \frac{\Gamma}{N} \\ & & \ddots & \vdots \\ & & & & \frac{\Gamma}{N} \end{bmatrix}$$
(10)

and $V_N \stackrel{\Delta}{=} [V_1, V_2, \dots, V_N]^t$. Using the distribution theory for transformations of random vectors [17], the joint pdf of V_1, V_2, \ldots, V_N can written as

$$f_{\boldsymbol{V}_{N}}\left(\left\{v_{n}\right\}_{n=1}^{N}\right) = Jf_{\boldsymbol{\gamma}_{[N]}}\left(\left\{\gamma_{[i]}\right\}_{i=1}^{N}\right)\Big|_{\boldsymbol{\gamma}_{[N]}=\boldsymbol{T}_{\mathrm{VB}}\boldsymbol{v}_{N}}$$
(11)

where J is the Jacobian of the virtual branch transformation and $\boldsymbol{v}_N \triangleq [v_1, v_2, \dots, v_N]^t.$ Denoting $I_N^{(i)}$ be the *i*th column of the $N \times N$ identity matrix

 I_N , we derive the recursion

$$\gamma_{[i]} = \left\langle I_N^{(i)}, \boldsymbol{\gamma}_{[N]} \right\rangle = \left\langle \boldsymbol{T}_{\mathrm{VB}}^t I_N^{(i)}, \boldsymbol{V}_N \right\rangle$$
$$= \underbrace{\left\langle \boldsymbol{T}_{\mathrm{VB}}^t I_N^{(i+1)}, \boldsymbol{V}_N \right\rangle}_{=\gamma_{i+1}} + \frac{\Gamma}{i} V_i \quad (12)$$

where $\gamma_{N+1} \stackrel{\Delta}{=} 0$ (or equivalently $I_N^{(N+1)} \stackrel{\Delta}{=} 0$) and $(\Gamma/i)V_i$ can be interpreted as the "difference between the adjacent ordered instantaneous SNR's." This implies that $0 < V_n < \infty$. Since the virtual branch transformation is linear and $T_{\rm VB}$ is an upper triangular matrix,

$$J = |\mathbf{T}_{\rm VB}| = \prod_{n=1}^{N} \frac{\Gamma}{n} = \frac{\Gamma^N}{N!}$$
(13)

where $|\cdot|$ denotes the determinant [23]. Note also that

$$\frac{1}{\Gamma} \left\langle \mathbf{1}_{N}, \boldsymbol{\gamma}_{[N]} \right\rangle = \frac{1}{\Gamma} \left\langle \boldsymbol{T}_{\mathrm{VB}}^{t} \mathbf{1}_{N}, \boldsymbol{V}_{N} \right\rangle = \left\langle \mathbf{1}_{N}, \boldsymbol{V}_{N} \right\rangle.$$
(14)

Substituting (13) and (14), in (11) together with (6), the joint pdf of V_1, V_2, \ldots, V_N becomes

$$f_{\boldsymbol{V}_N}\left(\{v_n\}_{n=1}^N\right) = \begin{cases} e^{-\langle \mathbf{1}_N, \boldsymbol{v}_N \rangle}, & 0 \le v_n < \infty\\ 0, & \text{otherwise.} \end{cases}$$
(15)

Therefore

$$f_{\mathbf{V}_{N}}\left(\{v_{n}\}_{n=1}^{N}\right) = \prod_{n=1}^{N} f_{V_{n}}(v_{n})$$
(16)

where $f_{V_n}(\cdot)$ is the pdf V_n given by

$$f_{V_n}(v) = \begin{cases} e^{-v}, & 0 \le v < \infty \\ 0, & \text{otherwise.} \end{cases}$$
(17)

This shows that the instantaneous SNR's of the virtual branches are i.i.d. with characteristic function (c.f.) given by

$$\psi_{V_n}(j\nu) \stackrel{\Delta}{=} \mathbb{E}\{e^{+j\nu V_n}\} = \frac{1}{1-j\nu}.$$
(18)

The instantaneous SNR of the combiner output can now be expressed in terms of the instantaneous SNR of the virtual branches as

$$\gamma_{\rm GDC} = \langle \boldsymbol{a}, \boldsymbol{T}_{\rm VB} \boldsymbol{V}_N \rangle. \tag{19}$$

Using the *independent* virtual branches, the N-fold nested integrals of (8) reduce to

$$P_{e,\text{GDC}} = \mathbb{E}_{\{V_n\}} \{ \Pr \{ e | \gamma_{\text{GDC}} = \langle \boldsymbol{a}, \boldsymbol{T}_{\text{VB}} \boldsymbol{V}_N \rangle \} \}$$
$$= \int_0^\infty \int_0^\infty \cdots \int_0^\infty \Pr \{ e | \langle \boldsymbol{T}_{\text{VB}}^t \boldsymbol{a}, \boldsymbol{v}_N \rangle \} \prod_{n=1}^N f_{V_n}(v_n) dv_n.$$
(20)

For many important modulation techniques, $\Pr\{e|\langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]}\rangle\}$ factors into a product of N terms, where each term depends only on one of the $\gamma_{[i]}$'s. Similarly, $\Pr\{e | \langle T_{VB}^t a, V_N \rangle\}$ factors into a product of N terms, each dependent only on one V_n . We will illustrate this by the following two important examples.

1) SEP for MPSK With GDC: For coherent detection of *M*-ary phase-shift keying (PSK), an alternative representation $\Pr\{e|\gamma_{GDC}\}$, involving a definite integral with *finite* limits, is given by [18], [19], [24]–[27]

$$\Pr\left\{e_{\text{MPSK}}|\gamma_{\text{GDC}}\right\} = \frac{1}{\pi} \int_{0}^{\Theta} \exp\left(-\frac{c_{\text{MPSK}}}{\sin^{2}\theta}\gamma_{\text{GDC}}\right) d\theta \quad (21)$$

where $c_{\text{MPSK}} = \sin^2(\pi/M)$ and $\Theta = \pi(M-1)/M$. Substituting (21) into (7), the SEP for MPSK becomes

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \mathbb{E}_{\gamma_{\text{GDC}}} \left\{ \exp\left(-\frac{c_{\text{MPSK}}}{\sin^{2}\theta}\gamma_{\text{GDC}}\right) \right\} d\theta$$
$$= \frac{1}{\pi} \int_{0}^{\Theta} \int_{0}^{\infty} \exp\left(-\frac{c_{\text{MPSK}}}{\sin^{2}\theta}\gamma\right) f_{\gamma_{\text{GDC}}}(\gamma) d\gamma d\theta.$$
(22)

Note that the inner integral of (22) is the c.f. of γ_{GDC} evaluated at $+j(c_{\text{MPSK}}/\sin^2\theta)$, and the SEP analysis that uses the c.f. of the combiner output SNR can be found in [19].

Although, the evaluation of (22) involves a single integration for averaging over the channel ensemble, it requires the knowledge of the pdf (or equivalently the c.f.) of γ_{GDC} . Alternatively, we substitute (21) into (8), and the SEP for MPSK becomes

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \mathbb{E}_{\{\gamma_{[i]}\}} \left\{ \exp\left(-\frac{c_{\text{MPSK}}}{\sin^{2}\theta} \langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]} \rangle\right) \right\} d\theta$$

$$= \frac{1}{\pi} \int_{0}^{\Theta} \int_{0}^{\infty} \int_{0}^{\gamma_{[1]}} \cdots \int_{0}^{\gamma_{[N-1]}} \exp\left(-\frac{c_{\text{MPSK}}}{\sin^{2}\theta} \langle \boldsymbol{a}, \boldsymbol{\gamma}_{[N]} \rangle\right)$$

$$\times f_{\boldsymbol{\gamma}_{[N]}} \left(\left\{ \gamma_{[i]} \right\}_{i=1}^{N} \right) d\gamma_{[N]} \cdots d\gamma_{[2]} d\gamma_{[1]} d\theta. \quad (24)$$

Note in (24) that, since the ordered physical branches are no longer independent, direct use of the methods given in [21] and [22] requires an N-fold nested integration for the expectation operation in (23). This is alleviated using the virtual branch technique by substituting (21) into (20) as

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \mathbb{E}_{\{V_n\}} \left\{ \exp\left(-\frac{c_{\text{MPSK}}}{\sin^2 \theta} \langle \boldsymbol{a}, \boldsymbol{T}_{\text{VB}} \boldsymbol{V}_N \rangle \right) \right\} d\theta$$
$$= \frac{1}{\pi} \int_{0}^{\Theta} \int_{0}^{\infty} \int_{0}^{\infty} \cdots \int_{0}^{\infty} \exp\left(-\frac{c_{\text{MPSK}}}{\sin^2 \theta} \langle \boldsymbol{T}_{\text{VB}}^t \boldsymbol{a}, \boldsymbol{v}_N \rangle \right)$$
$$\times \prod_{n=1}^{N} f_{V_n}(v_n) dv_n d\theta. \tag{25}$$

Exploiting the fact that V_n 's are independent, (25) becomes

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \prod_{n=1}^{N} \mathbb{E}_{V_n} \left\{ \exp\left(-\frac{c_{\text{MPSK}}b_n}{\sin^2\theta}V_n\right) \right\} d\theta$$
$$= \frac{1}{\pi} \int_{0}^{\Theta} \prod_{n=1}^{N} \psi_{V_n} \left(-\frac{c_{\text{MPSK}}b_n}{\sin^2\theta}\right) d\theta \tag{26}$$

where b_n is the *n*th element of $\boldsymbol{b} = \boldsymbol{T}_{VB}^t \boldsymbol{a}$. The powerfulness of the virtual branch technique is apparent by observing that the expectation operation in (23) no longer requires an *N*-fold nested integration.

Substituting (18) into (26) gives

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_0^{\Theta} \prod_{n=1}^N \left[\frac{\sin^2 \theta}{c_{\text{MPSK}} b_n + \sin^2 \theta} \right] d\theta.$$
(27)

Thus the derivation of the SEP for coherent detection of MPSK using N-branch GDC, involving the N-fold nested integrals in (24), essentially reduces to a single integral over θ with finite limits. The integrand is an N-fold product of a simple expression involving trigonometric functions. Note that the independence of the virtual branch variables plays a key role in simplifying the derivation.

2) SEP for MQAM With GDC: For coherent detection of squared M-ary quadrature amplitude modulation (QAM) with $M = 2^k$ for even k, $\Pr\{e|\gamma_{MRC}\}$ is given by [22]

$$\Pr\{e_{\text{MQAM}}|\gamma_{\text{GDC}}\} = q \frac{1}{\pi} \int_{0}^{\pi/2} \exp\left(-\frac{c_{\text{MQAM}}}{\sin^{2}\theta}\gamma_{\text{GDC}}\right) d\theta$$
$$-\frac{q^{2}}{4} \frac{1}{\pi} \int_{0}^{\pi/4} \exp\left(-\frac{c_{\text{MQAM}}}{\sin^{2}\theta}\gamma_{\text{GDC}}\right) d\theta$$
(28)

where $q = 4(1 - (1/\sqrt{M}))$, and $c_{MQAM} = 3/(2(M - 1))$. Using the virtual branch technique, similar to the steps used for MPSK, the SEP for MQAM becomes

$$P_{e,\text{GDC}}^{\text{MQAM}} = q \frac{1}{\pi} \int_0^{\pi/2} \prod_{n=1}^N \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} b_n + \sin^2 \theta} \right] d\theta$$
$$- \frac{q^2}{4} \frac{1}{\pi} \int_0^{\pi/4} \prod_{n=1}^N \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} b_n + \sin^2 \theta} \right] d\theta. \quad (29)$$

Again, the derivation of the SEP for coherent detection of MQAM using GDC in Rayleigh fading reduces to two terms, each consisting of a single integral over θ involving trigonometric functions with finite limits.

Note in passing that the two signal constellations for MQAM and MPSK coincide when M = 4, and hence their respective performance obtained from our virtual branch analysis are expected to be the same. Using the fact that the integrand in (29) is even symmetric about $\pi/2$ in θ , it is easy to verify that (29) with 4-QAM is identical to (27) 4-PSK.

III. APPLICATION OF GENERAL THEORY

The results given in (27) and (29) of Section II-C are for the SEP for coherent detection of MPSK and MQAM, respectively, using *N*-branch GDC in Rayleigh-fading channels. The b_n 's in (27) and (29) depend on the selection vector \boldsymbol{a} . The results obtained in (27) and (29) are *general* in the sense that they apply to a variety of diversity combining systems that fit the form of (2), including H-S/MRC, SD, and MRC. In the following, the general theory derived in Section II-C is used to evaluate the performance of H-S/MRC, SD, and MRC.

A. SEP's With H-S/MRC

The instantaneous output SNR of H-S/MRC is

$$\gamma_{\text{H-S/MRC}} = \sum_{i=1}^{L} \gamma_{[i]} \tag{30}$$

where $1 \le L \le N$. Note that $\gamma_{\text{H-S/MRC}} = \gamma_{\text{GDC}}$ with

$$\boldsymbol{a} = [\underbrace{1, 1, \dots 1}_{L \text{ terms}}, 0, \dots, 0]^t.$$
(31)

In this case,

$$b_n = \begin{cases} \Gamma, & n \le L \\ \Gamma \frac{L}{n}, & \text{otherwise.} \end{cases}$$
(32)

Substituting (32) into (27) of Section II-C, the SEP for MPSK with H-S/MRC can be easily obtained as

$$P_{e,\text{H-S/MRC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \left[\frac{\sin^{2} \theta}{c_{\text{MPSK}} \Gamma + \sin^{2} \theta} \right]^{L} \\ \times \prod_{n=L+1}^{N} \left[\frac{\sin^{2} \theta}{c_{\text{MPSK}} \Gamma \frac{L}{n} + \sin^{2} \theta} \right] d\theta. \quad (33)$$

Similarly, the SEP for MQAM with H-S/MRC can be obtained by substituting (32) into (29) of Section II-C as

$$P_{e,\mathrm{H}\text{-}\mathrm{S/MRC}}^{\mathrm{MQAM}} = q \frac{1}{\pi} \int_{0}^{\pi/2} \left[\frac{\sin^{2}\theta}{c_{\mathrm{MQAM}}\Gamma + \sin^{2}\theta} \right]^{L} \\ \times \prod_{n=L+1}^{N} \left[\frac{\sin^{2}\theta}{c_{\mathrm{MQAM}}\Gamma \frac{L}{n} + \sin^{2}\theta} \right] d\theta \\ - \frac{q^{2}}{4} \frac{1}{\pi} \int_{0}^{\pi/4} \left[\frac{\sin^{2}\theta}{c_{\mathrm{MQAM}}\Gamma + \sin^{2}\theta} \right]^{L} \\ \times \prod_{n=L+1}^{N} \left[\frac{\sin^{2}\theta}{c_{\mathrm{MQAM}}\Gamma \frac{L}{n} + \sin^{2}\theta} \right] d\theta. \quad (34)$$

B. SEP's With SD

SD is the simplest form of diversity system whereby the received signal from one of N diversity branches is selected [4]. The output SNR of SD is

$$\gamma_{\rm SD} \stackrel{\Delta}{=} \max_{i} \{\gamma_i\} = \gamma_{[1]}. \tag{35}$$

٦

Note that $\gamma_{\text{SD}} = \gamma_{\text{GDC}}$ with $\boldsymbol{a} = I_N^{(1)}$. In this case, $b_n = \Gamma/n$ for $n = 1, \ldots, N$, and substituting this into (27) and (29) of Section II-C, the SEP for coherent detection of MPSK and MQAM using SD becomes

$$P_{e,\text{SD}}^{\text{MPSK}} = \frac{1}{\pi} \int_0^{\Theta} \prod_{n=1}^N \left[\frac{\sin^2 \theta}{c_{\text{MPSK}} \Gamma \frac{1}{n} + \sin^2 \theta} \right] d\theta \qquad (36)$$

and

$$P_{e,\text{SD}}^{\text{MQAM}} = q \frac{1}{\pi} \int_{0}^{\pi/2} \prod_{n=1}^{N} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \Gamma \frac{1}{n} + \sin^2 \theta} \right] d\theta$$
$$- \frac{q^2}{4} \frac{1}{\pi} \int_{0}^{\pi/4} \prod_{n=1}^{N} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \Gamma \frac{1}{n} + \sin^2 \theta} \right] d\theta$$
(37)

respectively. Since SD is a special case of H-S/MRC with L =1, (36) and (37) can also be obtained from the H-S/MRC results of Section III-A by setting L = 1 in (33) and (34).

C. SEP's With MRC

In MRC, the received signals from all diversity branches are weighted and combined to maximize the SNR at the combiner output [4]. The output SNR of MRC is

$$\gamma_{\text{MRC}} \stackrel{\Delta}{=} \sum_{i=1}^{N} \gamma_i = \sum_{i=1}^{N} \gamma_{[i]}.$$
(38)

Note that $\gamma_{\text{MRC}} = \gamma_{\text{GDC}}$ with $\boldsymbol{a} = \mathbf{1}_N$. In this case, $b_n = \Gamma$ for $n = 1, \ldots, N$, and substituting this into (27) of Section II-C, the SEP for MPSK becomes

$$P_{e,\text{MRC}}^{\text{MPSK}} = \frac{1}{\pi} \int_{0}^{\Theta} \left[\frac{\sin^{2} \theta}{c_{\text{MPSK}} \Gamma + \sin^{2} \theta} \right]^{N} d\theta$$
$$\stackrel{\triangleq}{=} P_{e,\text{MRC}}^{\text{MPSK}}(N,\Gamma). \tag{39}$$

Specifically (39) is the SEP for coherent detection of MPSK using MRC with N independent branches having equal average SNR's of Γ in Rayleigh fading, and therefore a closed-form expression for (39) can be found in [28]. Comparing to the SEP expression given by [29, eq. (24)], we note that (39) is equivalent to the SEP for single-branch reception of MPSK in Nakagami-m fading with fading parameter N (i.e., m = N) having an average SNR of $N\Gamma$ [29]–[31].⁴

Similarly, the SEP for MQAM becomes

$$P_{e,\text{MRC}}^{\text{MQAM}} = q \frac{1}{\pi} \int_{0}^{\pi/2} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \Gamma + \sin^2 \theta} \right]^{N} d\theta$$
$$- \frac{q^2}{4} \frac{1}{\pi} \int_{0}^{\pi/4} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \Gamma + \sin^2 \theta} \right]^{N} d\theta$$
$$\triangleq P_{e,\text{MRC}}^{\text{MQAM}}(N,\Gamma). \tag{40}$$

The result of (40) is the SEP for coherent detection of MQAM using MRC with N independent branches having equal average SNR's of Γ in Rayleigh fading, and therefore a closed-form expression for (40) can be found in [32]. Note again that (40) is equivalent to the SEP for single-branch reception of MQAM in Nakagami fading with fading parameter N having an average SNR of NΓ [29]–[31].

Since MRC is a special case of H-S/MRC with L = N, (39) and (40) can also be obtained from the H-S/MRC results of Section III-A by setting L = N in (33) and (34).

IV. CANONICAL FORM FOR SEPS

A. Canonical Form for SEP's With GDC

The quest for obtaining insights from (27) and (29) is at its peak, which leads to an expansion for the integrand in (27). Let $\{\tilde{b}_n\}$ be the set of \tilde{N} distinct values of $\{b_n\}$ where each \tilde{b}_n has algebraic multiplicity μ_n such that $\sum_{n=1}^{\tilde{N}} \mu_n = N$. Then (27) and (29) can be rewritten as

$$P_{e,\text{GDC}}^{\text{MPSK}} = \frac{1}{\pi} \int_0^{\Theta} \prod_{n=1}^{\tilde{N}} \left[\frac{\sin^2 \theta}{c_{\text{MPSK}} \tilde{b}_n + \sin^2 \theta} \right]^{\mu_n} d\theta \quad (41)$$

and

$$P_{e,\text{GDC}}^{\text{MQAM}} = q \frac{1}{\pi} \int_{0}^{\pi/2} \prod_{n=1}^{\tilde{N}} \left[\frac{\sin^{2} \theta}{c_{\text{MQAM}} \tilde{b}_{n} + \sin^{2} \theta} \right]^{\mu_{n}} d\theta$$
$$- \frac{q^{2}}{4} \frac{1}{\pi} \int_{0}^{\pi/4} \prod_{n=1}^{\tilde{N}} \left[\frac{\sin^{2} \theta}{c_{\text{MQAM}} \tilde{b}_{n} + \sin^{2} \theta} \right]^{\mu_{n}} d\theta.$$
(42)

Letting $x = 1/\sin^2 \theta$ and $c_n = 1/c_{\text{MPSK}}\tilde{b}_n$, the integrand in (41) fits into the expression of (53) of the Appendix.⁵ Using the canonical expansion formula given in (54) of the Appendix, (41) can be rewritten as

$$P_{e,\text{GDC}}^{\text{MPSK}} = \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} \frac{1}{\pi} \int_0^{\Theta} \left[\frac{\sin^2 \theta}{c_{\text{MPSK}} \tilde{b}_n + \sin^2 \theta} \right]^k d\theta$$
(43)

where the weighting coefficients $A_{n,k}$ are given by (55). Comparing (43) with (39)

$$P_{e,\text{GDC}}^{\text{MPSK}} = \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} P_{e,\text{MRC}}^{\text{MPSK}}(k, \tilde{b}_n).$$
(44)

Interesting insights can now be obtained from (44). The SEP for MPSK using N-branch GDC in Rayleigh fading is simply

⁴This is due to the fact that pdf of $(2/\Gamma)\gamma_{MRC}$ is chi-squared distributed with 2N degrees of freedom, and, therefore, γ_{MRC} is equal (in distribution or law) to the square of the Nakagami random variable with m = N and mean $N\Gamma$. The fading parameter of Nakagami fading, usually denoted by the symbol m, is also known as fade parameter, fading severity factor, fading figure, or (inverse) fading-depth parameter.

⁵Although, there are numerous ways to expand the product of polynomials, we have chosen a specific one that leads to the canonical structure given by (44).

the weighted sum of the elementary SEP's. The weighting coefficients $A_{n,k}$ are given by (55), and the elementary SEP's for the (n, k)-entries are simply the SEP's for the coherent detection of MPSK using MRC with k independent branches having equal SNR's of \tilde{b}_n in Rayleigh-fading, or equivalently the SEP for single-branch reception of MPSK in Nakagami fading with fading parameter equal to k having an average SNR of $k\tilde{b}_n$.

Similarly, (42) can be rewritten as

$$P_{e,\text{GDC}}^{\text{MQAM}} = \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} q \frac{1}{\pi} \int_0^{\pi/2} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \tilde{b}_n + \sin^2 \theta} \right]^k d\theta \\ - \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} \frac{q^2}{4} \frac{1}{\pi} \int_0^{\pi/4} \left[\frac{\sin^2 \theta}{c_{\text{MQAM}} \tilde{b}_n + \sin^2 \theta} \right]^k d\theta.$$
(45)

Therefore,

$$P_{e,\text{GDC}}^{\text{MQAM}} = \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} P_{e,\text{MRC}}^{\text{MQAM}}(k, \tilde{b}_n).$$
(46)

Note that a similar structure, namely, "linear combination of the simple elementary SEP's," is evident from (46) for MQAM.

B. Canonical Forms for SEP's With H-S/MRC

For H-S/MRC, it can be seen from (32) that the number of distinct values of $\{b_n\}$ is $\tilde{N} = N - L + 1$. The distinct values of \tilde{b}_n s are given by

$$\tilde{b}_n = \begin{cases} \Gamma, & n = 1\\ \Gamma \frac{L}{L+n-1}, & n = 2, \dots, \tilde{N} \end{cases}$$
(47)

and their multiplicities μ_n s are given by

$$\mu_n = \begin{cases} L, & n = 1\\ 1, & n = 2, \dots, \tilde{N}. \end{cases}$$
(48)

Substituting (47) and (48) into (44) and (46) of Section IV-A, we arrive at the SEP's for coherent detection of MPSK and MQAM with H-S/MRC, which are given, respectively, by

$$P_{e,\text{H-S/MRC}}^{\text{MPSK}} = \sum_{k=1}^{L} A_{1,k} P_{e,\text{MRC}}^{\text{MPSK}}(k,\Gamma) + \sum_{n=2}^{N-L+1} A_{n,1} P_{e,\text{MRC}}^{\text{MPSK}}\left(1,\Gamma\frac{L}{L+n-1}\right)$$
(49)

$$P_{e,\text{H-S/MRC}}^{\text{MQAM}} = \sum_{k=1}^{L} A_{1,k} P_{e,\text{MRC}}^{\text{MQAM}}(k,\Gamma) + \sum_{n=2}^{N-L+1} A_{n,1} P_{e,\text{MRC}}^{\text{MQAM}}\left(1,\Gamma\frac{L}{L+n-1}\right).$$
(50)

The canonical structure for SEP with H-S/MRC is evident from (49) and (50) as a linear combination of the simple "elementary SEP's," as for the case of GDC in (44) and (46) of Section IV-A.

C. Canonical Forms for SEP's With SD

For SD, the number of distinct values of $\{b_n\}$, N, is equal to N. The distinct values are $\tilde{b}_n = \Gamma/n$ and the corresponding



Fig. 1. The symbol error probability for coherent detection of 8-PSK with H-S/MRC as a function of average SNR per branch in decibels for various L with N = 4. The curves are parameterized by different H-L/4 starting from the highest curve representing H-1/4, and decrease monotonically to the lowest curve representing H-4/4.



Fig. 2. The symbol error probability for coherent detection of 8-PSK with H-S/MRC as a function of average SNR per branch in decibels for various L with N = 8. The curves are parameterized by different H-L/8 starting from the highest curve representing H-1/8, and decrease monotonically to the lowest curve representing H-8/8.

multiplicities are given by $\mu_n = 1$ for n = 1, ..., N. Substituting these values into (44) and (46) of Section IV-A, the SEP's for coherent detection of MPSK and MQAM with SD becomes

$$P_{e,\text{SD}}^{\text{MPSK}} = \sum_{n=1}^{N} A_{n,1} P_{e,\text{MRC}}^{\text{MPSK}} \left(1, \frac{\Gamma}{n}\right)$$
(51)

and

$$P_{e,\text{SD}}^{\text{MQAM}} = \sum_{n=1}^{N} A_{n,1} P_{e,\text{MRC}}^{\text{MQAM}} \left(1, \frac{\Gamma}{n}\right)$$
(52)

respectively. Note again that the SEP for SD is simply a weighted sum of the elementary SEP's, as for the case of GDC in (44) and (46) of Section IV-A. Since SD is a special case



Fig. 3. The symbol error probability for coherent detection of 8-PSK with H-S/MRC as a function of average SNR per branch in decibels for various N with L = 2. The curves are parameterized by different H-2/N starting from the highest curve representing H-2/2, and decrease monotonically to the lowest curve representing H-2/8.



Fig. 4. The symbol error probability for coherent detection of 8-PSK with H-S/MRC as a function of average SNR per branch in decibels for various N with L = 4. The curves are parameterized by different H-4/N starting from the highest curve representing H-4/4, and decrease monotonically to the lowest curve representing H-4/8.

of H-S/MRC with L = 1, (51) and (52) can also be obtained alternatively from the H-S/MRC results given by (49) and (50) of Section IV-B by setting L = 1.

V. NUMERICAL EXAMPLES

In this section, the results derived in the previous section for H-S/MRC are illustrated. The notation H-L/N is used to denote H-S/MRC that selects and combines L out of N branches. Note that H-1/1 is a single branch receiver, and H-1/N and H-N/N are N-branch SD and MRC, respectively.

Figs. 1 and 2 show the SEP for coherent detection of MPSK with M = 8 (8-PSK) versus average SNR per branch for various L with N = 4 and N = 8, respectively. Note that SD and MRC



Fig. 5. The symbol error probability for coherent detection of 16-QAM with H-S/MRC as a function of average SNR per branch in decibels for various L with N = 4. The curves are parameterized by different H-L/4 starting from the highest curve representing H-1/4, and decrease monotonically to the lowest curve representing H-4/4.



Fig. 6. The symbol error probability for coherent detection of 16-QAM with H-S/MRC as a function of average SNR per branch in decibels for various L with N = 8. The curves are parameterized by different H-L/8 starting from the highest curve representing H-1/8, and decrease monotonically to the lowest curve representing H-8/8.

upper and lower bound, respectively, the SEP for H-S/MRC. It is seen that most of the gain of H-S/MRC is achieved for small L, e.g., H-S/MRC is within 1.1 dB of MRC when L = N/2.

Figs. 3 and 4 show the SEP for coherent detection of 8-PSK versus average SNR per branch for various N with L = 2 and L = 4, respectively. Note that, although the incremental gain with each additional antenna becomes smaller as N increases, the gain with each additional antenna is still significant even with N = 8. The results also show that, at a 10^{-4} SEP, H-2/8 requires 12.5 dB lower SNR than 2-branch MRC, and H-4/8 requires 4.5 dB lower SNR than 4-branch MRC.

Similar results for coherent detection of MQAM with M = 16 (16-QAM) are plotted in Figs. 5–8. These results show the same characteristics as 8-PSK illustrated in Figs. 1–4 except





that the SNR per branch is about 2 dB higher with 16-QAM to achieve the same SEP as 8-PSK.

VI. CONCLUSIONS

We derived exact expressions for the SEP for coherent detection of MPSK and MQAM with H-S/MRC in multipath-fading wireless environments. With H-S/MRC, L out of N diversity branches are selected and combined using MRC. This technique provides improved performance over L branch MRC when additional diversity is available, without requiring additional electronics and/or power. We considered independent Rayleigh fading on each diversity branch with equal signal-to-noise ratios, averaged over the fading. We analyzed this system using a novel "virtual branch" technique which resulted in a simple derivation of the SEP for *arbitrary* L and N. The key idea was to transform the dependent ordered-branch variables into a new set of i.i.d. *virtual branches*, and express the combiner output SNR as a linear combination of the i.i.d. virtual branch SNR variables. We further obtained a *canonical structure* for the SEP of H-S/MRC as a weighted sum of the elementary SEP's. The elementary SEP's are the SEP's using MRC with i.i.d. branches in Rayleigh fading, or equivalently the SEP's of the nondiversity (single-branch) system in Nakagami fading, whose closed-form expressions are well-known. Numerical results for 8-PSK and 16-QAM showed that H-S/MRC, even with $L \ll N$, can achieve performance close to that of *N*-branch MRC.

APPENDIX CANONICAL EXPANSION

The canonical expansion of

$$f(x) = \prod_{n=1}^{\tilde{N}} \left(\frac{c_n}{c_n + x}\right)^{\mu_n}$$
(53)

where the $\{-c_n\}$ are the \tilde{N} distinct poles of f(x), each having algebraic multiplicity μ_n , in terms of the elementary functions of the form $c_n/(c_n + x)$, is

$$f(x) = \sum_{n=1}^{\tilde{N}} \sum_{k=1}^{\mu_n} A_{n,k} \left(\frac{c_n}{c_n + x}\right)^k.$$
 (54)

The coefficients of the canonical expansion are given by

$$A_{n,k} = \frac{1}{c_n^k (\mu_n - k)!} f_n^{(\mu_n - k)}(0),$$

$$n = 1, \dots, \tilde{N}, \quad k = 1, \dots, \mu_n \quad (55)$$

where $f_n^{(k)}(0)$ denotes the kth derivative of $f_n(x) = x^{\mu_n} f(x - c_n)$ evaluated at x = 0.

ACKNOWLEDGMENT

The authors wish to thank G. J. Foschini, L. A. Shepp, N. C. Beaulieu, D. P. Taylor, J. G. Proakis, P. F. Dahm, and M. Shtaif for helpful discussions.

REFERENCES

- J. H. Winters, J. Salz, and R. Gitlin, "The impact of antenna diversity on the capacity of wireless communication system," *IEEE Trans. Commun.*, vol. 42, pp. 1740–1751, Feb./Mar./Apr. 1994.
- [2] G.-T. Chyi, J. G. Proakis, and K. M. Keller, "On the symbol error probability of maximum-selection diversity reception schemes over a Rayleigh fading channel," *IEEE Trans. Commun.*, vol. 37, pp. 79–83, Jan. 1989.
- [3] E. A. Neasmith and N. C. Beaulieu, "New results on selection diversity," *IEEE Trans. Commun.*, vol. 46, pp. 695–704, May 1998.
- [4] W. C. Jake, Ed., Microwave Mobile Communications, (IEEE press classic reissue) ed. Piscataway, NJ: IEEE Press, 1995.
- [5] J. H. Winters, "Smart antennas for wireless systems," *IEEE Pers. Commun. Mag.*, pp. 23–27, Feb. 1998.
- [6] H. Erben, S. Zeisberg, and H. Nuszkowski, "BER performance of a hybrid SC/MRC 2DPSK RAKE receiver in realistic mobile channels," in *Proc. 44th Annu. Int. Vehicular Technology Conf.*, vol. 2, Stockholm, Sweden, June 1994, pp. 738–741.





- [7] T. Eng, N. Kong, and L. B. Milstein, "Comparison of diversity combining techniques for Rayleigh-fading channels," *IEEE Trans. Commun.*, vol. 44, pp. 1117–1129, Sept. 1996.
- [8] N. Kong and L. B. Milstein, "Combined average SNR of a generalized diversity selection combining scheme," in *Proc. IEEE Int. Conf.* on Communications, vol. 3, Atlanta, GA, June 1998, pp. 1556–1560.
- [9] M. Z. Win and J. H. Winters, "Analysis of hybrid selection/maximalratio combining of diversity branches with unequal SNR in Rayleigh fading," in *Proc. 49th Annu. Int. Vehicular Technology Conf.*, vol. 1, Houston, TX, May 1999, pp. 215–220.
- [10] —, "Analysis of hybrid selection/maximal-ratio combining in Rayleigh fading," in *Proc. IEEE Int. Conf. on Communications*, vol. 1, Vancouver, Canada, June 1999, pp. 6–10.
- [11] —, "Analysis of hybrid selection/maximal-ratio combining in Rayleigh fading," *IEEE Trans. Commun.*, vol. 47, pp. 1773–1776, Dec. 1999.
- [12] M.-S. Alouini and M. K. Simon, "Performance analysis of generalized selective combining over Rayleigh fading channels," in *Proc. 8th Communications Theory Mini Conf.*, Vancouver, Canada, June 1999, pp. 110–114.
- [13] A. W. Naylor and G. R. Sell, *Linear Operator Theory in Engineering and Science*, 2nd ed. New York: Springer-Verlag, 1982.
- [14] J. B. Conway, *A Course in Functional Analysis*, 2nd ed. New York: Springer-Verlag, 1990.
- [15] A. N. Shiryaev, Probability, 2nd ed. New York: Springer-Verlag, 1995.
- [16] R. Durrett, Probability: Theory and Examples, first ed. Pacific Grove, CA: Wadsworth & Brooks/Cole, 1991.
- [17] P. J. Bickel and K. Doksum, *Mathematical Statistics: Basic Ideas and Selected Topics*, first ed. Oakland, CA: Holden-Day, 1977.
- [18] C. Tellambura, A. J. Mueller, and V. K. Bhargava, "Analysis of M-ary phase-shift keying with diversity reception for land-mobile satellite channels," *IEEE Trans. Veh. Technol.*, vol. 46, pp. 910–922, Nov. 1997.
- [19] A. Annamalai, C. Tellambura, and V. K. Bhargava, "A unified approach to performance evaluation of diversity systems on fading channels," in *Wireless Multimedia Network Technologies*, R. Ganesh and Z. Zvonar, Eds. Amsterdam, The Netherlands: Kluwer Academic, 1999, pp. 311–330.
- [20] —, "Exact evaluation of maximal-ratio and equal-gain diversity receivers for *M*-ary QAM on Nakagami fading channels," *IEEE Trans. Commun.*, vol. 47, pp. 1335–1344, Sept. 1999.
- [21] M. K. Simon and D. Divsalar, "Some new twists to problems involving the Gaussian probability integral," *IEEE Trans. Commun.*, vol. 46, pp. 200–210, Feb. 1998.
- [22] M.-S. Alouini and A. Goldsmith, "A unified approach for calculating error rates of linearly modulated signals over generalized fading channels," in *Proc. IEEE Int. Conf. on Communications*, vol. 1, Atlanta, GA, June 1998, pp. 459–463.
- [23] P. D. Lax, Linear Algebra, 1st ed. New York: Wiley, 1996.
- [24] F. S. Weinstein, "Simplified relationship for the probability distribution of the phase of a sine wave in narrow-band normal noise," *IEEE Trans. Inform. Theory*, vol. IT-20, pp. 658–661, Sept. 1974.
- [25] R. F. Pawula, S. O. Rice, and J. H. Roberts, "Distribution of the phase angle between two vectors perturbed by Gaussian noise," *IEEE Trans. Commun.*, vol. COM-30, pp. 1828–1841, Aug. 1982.
- [26] J. W. Craig, "A new, simple and exact result for calculating the probability of error for two-dimensional signal constellations," in *Proc. Military Communications Conf.*, Boston, MA, 1991, pp. 25.5.1–25.5.5.
- [27] M. K. Simon, S. M. Hinedi, and W. C. Lindsey, *Digital Communication Techniques: Signal Design and Detection*, first ed. Englewood Cliffs, NJ: Prentice Hall, 1995.
- [28] S. Chennakeshu and J. B. Anderson, "Error rates for Rayleigh fading multichannel reception of MPSK signals," *IEEE Trans. Commun.*, vol. 43, pp. 338–346, Feb./Mar./Apr. 1995.
- [29] M. Z. Win and J. H. Winters, "Exact error probability expressions for MRC in correlated Nakagami channels with unequal fading parameters and branch powers," in *Proc. IEEE Global Telecomm. Conf., Symp. on Communications Theory*, vol. 1, Rio de Janeiro, Brazil, Dec. 1999, pp. 2331–2335.

- [30] —, "On maximal ratio combining in correlated Nakagami channels with unequal fading parameters and SNR's among branches: An analytical framework," in *Proc. IEEE Wireless Communications and Networking Conf.*, vol. 3, New Orleans, LA, Sept. 1999, pp. 1058–1064.
- [31] M. Z. Win, G. Chrisikos, and J. H. Winters, "MRC performance for M-ary modulation in arbitrarily correlated Nakagami fading channels," *IEEE Commun. Lett.*, vol. 4, pp. 301–303, Oct. 2000.
- [32] J. Lu, T. T. Tjhung, and C. C. Chai, "Error probability performance of *L*-branch diversity reception of MQAM in Rayleigh fading," *IEEE Trans. Commun.*, vol. 46, pp. 179–181, Feb. 1998.



Moe Z. Win (S'85–M'87–SM'97) received the B.S. degree (*magna cum laude*) from Texas A&M University, College Station, and the M.S. degree from the University of Southern California (USC), Los Angeles, in 1987 and 1989, respectively, both in electrical engineering. As a Presidential Fellow at USC, he received both an M.S. degree in applied mathematics and the Ph.D. degree in electrical engineering in 1998.

In 1987, he joined the Jet Propulsion Laboratory (JPL), California Institute of Technology, Pasadena.

From 1994 to 1997, he was a Research Assistant with the Communication Sciences Institute at USC, where he played a key role in the successful creation of the Ultra-Wideband Radio Laboratory. Since 1998, he has been with the Wireless Systems Research Department, AT&T Laboratories-Research, Middletown, NJ, where he is a Principal Technical Staff Member. His main research interests are the application of communication, detection, and estimation theories to a variety of communications problems including time-varying channels, diversity, equalization, synchronization, signal design, ultrawide-bandwidth communication, and optical communications.

Dr. Win is a member of Eta Kappa Nu, Tau Beta Pi, Pi Mu Epsilon, Phi Theta Kappa, and Phi Kappa Phi. He was a University Undergraduate Fellow at Texas A&M University, where he received, among others awards, the Academic Excellence Award. At USC, he received several awards including the Outstanding Research Paper Award and the Phi Kappa Phi Student Recognition Award. He was the recipient of the IEEE Communications Society Best Student Paper Award at the Fourth Annual IEEE NetWorld+Interop '97 Conference. He has been involved actively in chairing and organizing sessions and has served as a member of the Technical Program Committee in a number of IEEE conferences. He currently serves as the Technical Program Chair for the IEEE Communication Theory Symposium of ICC-2003 and IEEE Conference on Ultra Wideband Systems and Technologies (2002), and Technical Program Vice-Chair for IEEE International Conference on Communications (2002). He served as the Tutorial Chair for IEEE Semiannual International Vehicular Technology Conference (Fall-2001) and the Technical Program Chair for the IEEE Communication Theory Symposium of Globecom-2000. He is the current Editor for Equalization and Diversity for the IEEE TRANSACTIONS ON COMMUNICATIONS and a Guest-Editor for the 2002 IEEE JOURNAL ON SELECTED AREAS IN COMMUNICATIONS, Special Issue on ultra wide band radio in multi-access wireless communications.



Jack H. Winters (S'77–M'81–SM'88–F'96) received the B.S.E.E. degree from the University of Cincinnati, Cincinnati, OH, in 1977, and the M.S. and Ph.D. degrees in electrical engineering from The Ohio State University, Columbus, in 1978 and 1981, respectively.

Since 1981, he has been with AT&T Bell Laboratories, and now AT&T Labs-Research, where he is currently Division Manager of the Wireless Systems Research Department. He has studied signal processing techniques for increasing the capacity and reducing

signal distortion in fiber optic, mobile radio, and indoor radio systems and is currently studying smart antennas, adaptive arrays, and equalization for wireless local area networks and mobile radio systems.